

## Applied Natural Language Processing

Info 256
Lecture 3: Finding Distinctive Terms (Aug 30, 2023)
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## Panel B: Phrases Used More Often by Republicans

| Two-Word Phrases |  |  |
| :--- | :--- | :--- |
| stem cell | personal accounts | retirement accounts |
| natural gas | Saddam Hussein | government spending |
| death tax | pass the bill | national forest |
| illegal aliens | private property | urge support |
| class action | border security | cell lines |
| war on terror | President announces | cord blood |
| embryonic stem | human life | action lawsuits |
| tax relief | Chief Justice | economic growth |
| illegal immigration | human embryos | food program |
| date the time | increase taxes |  |
| Three-Word Phrases |  | Tongass national forest |
| embryonic stem cell | Circuit Court of Appeals | pluripotent stem cells |
| hate crimes legislation | death tax repeal | Supreme Court of Texas |
| adult stem cells | housing and urban affairs | Justice Priscilla Owen |
| oil for food program <br> personal retirement accounts <br> energy and natural resources <br> global war on terror | million jobs created <br> hate crimal flood insurance <br> oil for food scandal | Justice Janice Rogers |
| change hearts and minds | private property rights | American Bar Association |
| global war on terrorism | temporary worker program | growth and job creation |
| class action reform | natural gas natural |  |



Which are the words most likely to be from Android and most likely from iPhone?


[^0]
## Distinctive terms

- Finding distinctive terms is useful:
- As a data exploration exercise to understand larger trends in individual word differences).
- As a pre-processing step of feature selection.
- When the two datasets are A and $\neg \mathrm{A}$, these terms also provide insight into what $A$ is about.
- Many methods for finding these terms! (Developed in NLP, corpus linguistics, political science, etc.)


## Difference in proportions

For word w written by author with label k (e.g., \{democrat, republican\}), define the frequency to be the normalized count of that word

$$
f_{w, k}=\frac{C(w, k)}{\sum_{w^{\prime}} C\left(w^{\prime}, k\right)}
$$

count of word w in group k
count of all words in group k

$$
f_{w, k=\text { dem }}-f_{w, k=\text { repub }}
$$

Partisan Words, 106th Congress, Abortion
(Difference of Proportions)


## Difference in proportions

- The difference in proportions is a conceptually simple measure and easily interpretable.
- Drawback: tends to emphasize words with high frequency (where even comparatively small differences in word usage between groups is amplified).
- Also, no measure whether a difference is statistically meaningful. We have uncertainty about the what the true proportion is for any group.


## $x^{2}$

- $X^{2}$ (chi-square) is a statistical test of dependence--here, dependence between the two variables of word identity and corpus identity.
- For assessing the difference in two datasets, this test assumes a $2 x 2$ contingency table:

|  | word | $\neg$ word |
| :---: | :---: | :---: |
| corpus 1 | 7 | 104023 |
| corpus 2 | 104 | 251093 |
|  |  |  |

## $x^{2}$

Does the word robot occur significantly more frequently in science fiction?

|  | robot | ᄀrobot |
| ---: | :---: | :---: |
|  | 104 | 1004 |
| sci-fi | $=10.3 \%$ |  |
|  | 2 | 13402 |
|  |  | $=0.015 \%$ |

## $x^{2}$

For each cell in contingency table, sum the squared difference between observed value in cell and the expected value assuming independence.

$$
\chi^{2}=\sum_{i, j} \frac{\left(O_{i j}-E_{i j}\right)^{2}}{E_{i j}}
$$

| sci-fi | robot | $\neg$ robot | sum | frequency |
| :---: | :---: | :---: | :---: | :---: |
|  | 104 | 1004 | 1108 | 0.076 |
| $\neg \mathrm{SCi}-\mathrm{fi}$ | 2 | 13402 | 13404 | 0.924 |
| sum | 106 | 14406 |  |  |
| frequency | 0.007 | 0.993 |  |  |

$$
\begin{aligned}
P(\text { robot, scifi }) & =P(\text { robot }) \times P(\text { scifi }) \\
& =0.007 \times 0.076=0.00053
\end{aligned}
$$

Among 14512 words, we would expect to see 7.69 occurrences of robot in sci-fi texts.

| sci-fi | robot | ᄀrobot | $P($ scifi $)$ | 0.076 |
| :---: | :---: | :---: | :---: | :---: |
|  | 7.69 | 1095.2 |  |  |
| $\neg$ Sci-fi | 93.9 | 13315.2 | $P(\neg \mathrm{scifi})$ | 0.924 |

$P$ (robot) $P$ ( $\neg$ robot)

| 0.007 | 0.993 |
| :--- | :--- |

## $x^{2}$

- What $x^{2}$ is asking is: how different are the observed counts from the counts we would expect given complete independence?

|  | robot | $\neg$ robot |
| :---: | :---: | :---: |
|  | 104 | 1004 |
|  | sci-fi-fi | 2 |
|  | 13402 |  |
|  |  |  |


|  | robot | $\neg$ robot |
| :---: | :---: | :---: |
|  | 7.69 | 1095.2 |
|  | sci-fi |  |
|  | 93.9 | 13315.2 |
|  |  |  |

## $x^{2}$

- With algebraic manipulation, simpler form for $2 \times 2$ table O (cf. Manning and Schütze 1999)

$$
\chi^{2}=\frac{N\left(O_{11} O_{22}-O_{12} O_{21}\right)^{2}}{\left(O_{11}+O_{12}\right)\left(O_{11}+O_{21}\right)\left(O_{12}+O_{22}\right)\left(O_{21}+O_{22}\right)}
$$

## $x^{2}$

- The $x^{2}$ value is a statistic of dependence with a probability governed by a $x^{2}$ distribution; if this value has low enough probability in that measure, we can reject the null hypothesis of the independence between the two variables.


## $x^{2}$



## $x^{2}$

- Chi-square is ubiquitous in corpus linguistics (and in NLP as a measure of collocations).
- A few caveats for its use:
- Each cell should have an expected count of at least 5
- Each observation is independent


## $x^{2}$

- A drawback, however, is due to the burstiness of language: the tendency for the same words to clump together in texts.
- Chi-square is testing for independence of two variables (word identity and corpus identity), but it assumes each mention of the word is independent from the others.

- Is Dracula really a word that distinguishes these two corpora?
- It distinguishes one text, but otherwise doesn't appear in the corpus at all.


## Mann-Whitney rank sums test

- Mann-Whitney is a test of the difference in some quantity of interest in two datasets. Null hypothesis: if you select a random sample from group A and another from group B , just as likely that A will be greater than $B$ as less than $B$.

| A | A | A | A | A | A | A | A |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 2 | 1 | 4 | 3 | 2 | 0 | 1 |


| B | B | B | B | B | B |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 8 | 4 | 9 | 7 | 6 | 10 |

## Mann-Whitney

| A | A | A | A | A | A | A | A | B | B | B | B | B | B | B |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 1 | 4 | 3 | 2 | 0 | 1 | 8 | 4 | 9 | 7 | 2 | 10 | 5 |


| A | A | A | A | A | A | B | A | B | B | B | B | B | B |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 1 | 1 | 1 | 2 | 2 | 2 | 3 | 4 | 5 | 7 | 8 | 9 | 10 |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 |

## Mann-Whitney

|  | A | A | A | A | A | A | A | A | B | B | B | B | B | B | B |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 1 | 4 | 3 | 2 | 0 | 1 | 8 | 4 | 9 | 7 | 2 | 10 | 5 |
|  | A | A | A | A | A | A | B | A | B | B | B | B | B | B | B |
|  | 0 | 1 | 1 | 1 | 2 | 2 | 2 | 3 | 4 | 5 | 7 | 8 | 9 | 9 | 10 |
| ranks | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 3 | 14 |

## Mann-Whitney

$$
\begin{gathered}
\mathrm{R}_{1}=7+9+10+11+12+13+14=76 \\
U_{1}=R_{1}-\frac{n_{1}\left(n_{1}+1\right)}{2}
\end{gathered}
$$

$n_{1}=$ sample size for dataset from which $R_{1}$ is derived (e.g., number of chunks)

- Once we have this $U$ value, we can ask whether it's significantly different from the average value we would expect if there's no difference between the two groups at all. (We can do so by converting $U$ to a z-score using a Normal approximation and checking significance).

$$
\begin{array}{cccccccc}
\text { A } & \text { A } & \text { A } & \text { A } & \text { A } & \text { A } & \text { A } & \text { A } \\
1 & 2 & 1 & 4 & 3 & 2 & 0 & 1
\end{array} \quad \begin{array}{cccccccc}
\text { B } & \text { B } & \text { B } & \text { B } & \text { B } & \text { B } & \text { B } \\
8 & 4 & 9 & 7 & 6 & 10 & 5
\end{array}
$$

- In corpus linguistics, each measurement is the count of a word in a fixedsized chunk of text (e.g., 500 words).
- This lets us accommodate a more realistic assumption about the burstiness of language.

$$
\begin{aligned}
& \begin{array}{lllclccc|cccccc}
\text { A } & \text { A } & \text { A } & \text { A } & \text { A } & \text { A } & \text { A } & \text { A } & \text { B } & \text { B } & \text { B } & \text { B } & \text { B } & \text { B } \\
0 & 0 & 0 & 417 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0
\end{array} \\
& \text { Not a significant } \\
& \text { difference in ranks }
\end{aligned}
$$

## Log-odds ratios with priors

- The odds of a word is another informative measure:

$$
\begin{aligned}
& \frac{\text { probability of event occuring in corpus } X}{\text { probability of event not occuring in corpus } X} \\
& \frac{\text { count of word } w \text { in corpus } X}{\# \text { words in } X} \\
& \frac{\text { unt of all words } \neg w \text { in corpus } X}{\# \text { words in } X}=\frac{\text { count of word } w \text { in corpus } X}{\text { count of all words } \neg w \text { in corpus } X}
\end{aligned}
$$

$$
\frac{\# \text { mentions of "recall" }}{\# \text { mentions of } \neg \text { "recall" }}
$$

$$
\frac{14}{1,000,000}=0.000014
$$

## Log-odds ratios with priors

- Two get a measure of difference, we can compare two odds in different corpora

```
Democrat n=1000014
```


## Republican $\mathrm{n}=2000042$

$$
\frac{\text { \# mentions of "recall" }}{\text { \# mentions of } ᄀ \text { "recall" }}
$$

$\frac{42}{2,000,000}=0.000021$

- The odds ratio gives us one way of combining these into a single score

$$
\frac{\frac{14}{1,000,000}}{\frac{42}{2,000,000}}=0.667
$$

## Log-odds ratios with priors

- But this is bounded by $(0, \infty)$ and not easy to interpret with respect to the boundary (1) separating a word being likelier in corpus than another.
- We can work with the log instead, which transforms this into the space $(-\infty, \infty)$, with 0 as a boundary

$$
\frac{\frac{14}{1,000,000}}{\frac{42}{2,000,000}}=0.667
$$

$$
\log \left(\frac{\frac{14}{1,000,000}}{\frac{42}{2,000,000}}\right)=-0.4054
$$

## Log-odds ratios with priors

- What if we have 0 counts?

$$
\log \left(\frac{\frac{14}{1,000,000}}{\frac{0}{2,000,000}}\right)=
$$

- We can add pseudocounts! e.g., assume vocabulary size of 10,000 words, 100 here $=10,000$ * 0.01 to account for total pseudocount mass added, and we remove 0.01 from the denominators since

$$
\log \left(\frac{\frac{14+0.01}{\frac{1,000,000+100-0.01}{0+0.01}}}{\frac{0,000,000+100-0.01}{}}\right)=7.94
$$ the denominator is the count of $\neg$ word.

## Log-odds ratios with priors

$$
\log \left(\frac{\frac{14+0.01}{\frac{1,000,000+100-0.01}{42+0.01}}}{\frac{4,000,000+100-0.01}{2}}\right)=-0.4050
$$

$$
=\log \left(\frac{14+0.01}{1,000,000+100-0.01}\right)-\log \left(\frac{42+0.01}{2,000,000+100-0.01}\right)
$$

## Log-odds ratios with priors

- Transform them into z-scores by dividing them by the standard deviation.

$$
\approx \frac{\log \left(\frac{14+0.01}{1,000,000+100-0.01}\right)-\log \left(\frac{42+0.01}{2,000,000+100-0.01}\right)}{\sqrt{\frac{1}{14+0.01}+\frac{1}{42+0.01}}}
$$

The larger the term counts (e.g., 14, 42), the more confident we can be that the difference is meaningful

## Other methods

- There are many other methods for learning distinguishing words between two corpus; major classes:
- Model-based methods that assume parametric forms + Bayesian priors (for smoothing) [Monroe et al. 2009]
- Methods using classification to learn informative features that separate classes.


## Activity

- Hypothesize terms that will be different between 2020 Democrat and Republican platforms.
- Execute chi-square to find terms that are different
- Compare to Mann-Whitney for this data.


[^0]:    http://varianceexplained.org/r/trump-tweets/

